

Q) Random sample of equal sizes i.e., $n_1 = n_2 = n$ drawn from two independent normal populations $N(\mu_1, \sigma_1^2)$ and $N(\mu_2, \sigma_2^2)$, where μ_1 and μ_2 both are unknown also σ_1^2 and σ_2^2 are also unknown obtain $100(1-\alpha)\%$ confidence interval for the parametric function $\psi = \mu_1/\mu_2$.

Proof

Consider the random variables $U = X_1 - \psi(X_2)$, where X_1 is a $N(\mu_1, \sigma_1^2)$ and X_2 is another $N(\mu_2, \sigma_2^2)$ variate. Then expectation of U ,

$$E(U) = E(X_1) - \psi E(X_2)$$

$$= \mu_1 - \frac{\mu_1}{\mu_2} \mu_2$$

$$= 0$$

$$\text{and } V(U) = V(X_1) + \psi^2 V(X_2) - 2\psi \text{cov}(X_1, X_2)$$

$$= \sigma_1^2 + \frac{\mu_1^2}{\mu_2^2} \sigma_2^2$$

$$= V(X_1) + \psi^2 V(X_2)$$

[$\because X_1$ and X_2 are independent]

$$\sigma_u^2 = \sigma_1^2 + \psi^2 \sigma_2^2$$

$$\hat{\sigma}_u^2 = s_u^2 = s_1^2 + \psi^2 s_2^2$$

$$\text{where, } s_1^2 = \frac{1}{n_1 - 1} \sum (\alpha_{1i} - \bar{\alpha}_1)^2$$

$$\text{and } s_2^2 = \frac{1}{n_2 - 1} \sum (\alpha_{2j} - \bar{\alpha}_2)^2$$

$$\therefore \bar{U} = \bar{X}_1 - \gamma \bar{X}_2$$

$$E(\bar{U}) = E(\bar{X}_1) - \gamma E(\bar{X}_2)$$

$$= \mu_1 - \frac{\mu_1}{\mu_2} \cdot \mu_2$$

$$= 0$$

$$V(\bar{U}) = V(\bar{X}_1) + \gamma^2 V(\bar{X}_2)$$

$$= \frac{\sigma_1^2}{n} + \gamma^2 \frac{\sigma_2^2}{n}$$

$$= \frac{\sigma_1^2 + \gamma^2 \sigma_2^2}{n}$$

$$\therefore \bar{U} = \bar{X}_1 - \gamma \bar{X}_2 \text{ is a } N\left(0, \frac{\sigma_1^2 + \gamma^2 \sigma_2^2}{n}\right)$$

Since σ_1^2 and σ_2^2 are not known, the statistic

$$g(T, \theta) = \frac{\bar{x}_1 - \gamma \bar{x}_2}{s_u/\sqrt{n}} \sim t \text{ with } (n-1) \text{ d.f.}$$

$$\therefore P_{\theta} \left[\left| \frac{\bar{x}_1 - \gamma \bar{x}_2}{s_u/\sqrt{n}} \right| \leq t_{\alpha/2, n-1} \right] = 1 - \alpha$$

$$\text{or } P_{\theta} \left[(\bar{x}_1 - \gamma \bar{x}_2)^2 \leq t_{\alpha/2, n-1}^2 \cdot \frac{s_u^2}{n} \right] = 1 - \alpha$$

$$\text{or } P_{\theta} \left[(\bar{x}_1 - \gamma \bar{x}_2)^2 \leq t_{\alpha/2, n-1}^2 \cdot \frac{s_1^2 + \gamma^2 s_2^2}{n} \right] = 1 - \alpha$$

$$0.14 P_0 \left[\bar{x}_1^2 - 2\gamma \bar{x}_1 \bar{x}_2 + \gamma^2 \bar{x}_2^2 \leq t_{\alpha/2, n-1}^2 \frac{s_1^2 + \gamma^2 s_2^2}{n} \right] = 1 - \alpha$$

$$0.14 P_0 \left[\bar{x}_1^2 - 2\gamma \bar{x}_1 \bar{x}_2 + \gamma^2 \bar{x}_2^2 - t_{\alpha/2, n-1}^2 \frac{s_1^2}{n} - t_{\alpha/2, n-1}^2 \frac{\gamma^2 s_2^2}{n} \leq 0 \right] = 1 - \alpha$$

$$0.14 P_0 \left[\gamma^2 \left(\bar{x}_2^2 - \frac{t_{\alpha/2, n-1}^2 s_2^2}{n} \right) - 2\gamma \bar{x}_1 \bar{x}_2 + \bar{x}_1^2 - t_{\alpha/2, n-1}^2 \frac{s_1^2}{n} \leq 0 \right] = 1 - \alpha$$

If θ_1 and θ_2 be the roots ($\theta_2 > \theta_1$) of the quadratic equation in γ ,

$$\gamma^2 \left(\bar{x}_2^2 - \frac{t_{\alpha/2, n-1}^2 s_2^2}{n} \right) - 2\gamma \bar{x}_1 \bar{x}_2 + \bar{x}_1^2 - t_{\alpha/2, n-1}^2 \frac{s_1^2}{n} = 0$$

then $100(1-\alpha)\%$ confidence limits of $\gamma = \mu_1/\mu_2$ will be between θ_1 and θ_2 .

Confidence Interval of Population Proportion

Confidence Interval of single proportion

Let, x_1, x_2, \dots, x_n be the random samples of size n drawn from the Bernoulli population with parameter p .

Let, x be the statistic (= no. of units belonging to a particular characteristic) defined for the purpose.

$$\therefore x \sim B(n, p)$$

Now, $E(x) = np$ and $V(x) = np(1-p)$

Then, x/n is the sample proportion of that characteristic = p (say)

Therefore, $E(p) = E(x/n) = p$

$$\begin{aligned} \text{and } V(p) &= 1/n^2 V(x) \\ &= \frac{p(1-p)}{n} \end{aligned}$$

Under large sample normality assumption

$$p \sim N\left(p, \frac{p(1-p)}{n}\right),$$

$$\therefore \frac{p - \hat{p}}{\sqrt{\frac{p(1-p)}{n}}} \sim N(0, 1)$$

$$\therefore \text{We have } P_0 \left[-z_{\alpha/2} \leq \frac{p - \hat{p}}{\sqrt{\frac{p(1-p)}{n}}} \leq +z_{\alpha/2} \right] = 1 - \alpha$$

$$\text{or } P_0 \left[-z_{\alpha/2} \cdot \sqrt{\frac{\hat{p}(1-\hat{p})}{n}} \leq p - \hat{p} \leq +z_{\alpha/2} \cdot \sqrt{\frac{\hat{p}(1-\hat{p})}{n}} \right] = 1 - \alpha$$

$$\text{or } P_0 \left[- \left(\hat{p} + z_{\alpha/2} \cdot \sqrt{\frac{\hat{p}(1-\hat{p})}{n}} \right) \leq p \leq - \left(\hat{p} - z_{\alpha/2} \cdot \sqrt{\frac{\hat{p}(1-\hat{p})}{n}} \right) \right] = 1 - \alpha$$

$$\text{or } P_0 \left[\hat{p} - z_{\alpha/2} \cdot \sqrt{\frac{\hat{p}(1-\hat{p})}{n}} \leq p \leq \hat{p} + z_{\alpha/2} \cdot \sqrt{\frac{\hat{p}(1-\hat{p})}{n}} \right] = 1 - \alpha$$

Hence $100(1-\alpha)\%$ confidence limits for

$$p \text{ will be } \hat{p} \pm z_{\alpha/2} \cdot \sqrt{\frac{\hat{p}(1-\hat{p})}{n}} = \hat{p} \pm z_{\alpha/2} \cdot \sqrt{\frac{p(1-p)}{n}}$$